

# MULTIVARIATE CONDITIONAL & MARGINAL NORMAL DISTRIBUTION: A PROBABILITY APPROACH

**Zahid Iqbal\***, **Junaid Saghir Siddiqui\*\***

\* Dept. Mathematics and Statistics, International Islamic University, Islamabad  
 Email.: [Z.iqbal@iit.edu.pk](mailto:Z.iqbal@iit.edu.pk). Mobile:+92-0300-2125532, office: 092-051-9019735  
 \*\*Dept. Statistics, University of Karachi

**ABSTRACT:** Suppose that vector random variable  $X = (X_1, X_2, X_3, \dots, X_p)^T$  has a multivariate normal distribution then the probability distribution function (PdF) is given by

$$f(X) = \frac{e^{-\frac{1}{2}[X-\mu]\Sigma^{-1}[X-\mu]}}{(2\pi)^{\frac{p}{2}} |\Sigma|^{\frac{1}{2}}}$$

where variance-covariance matrix  $=\Sigma$  and is of order  $p \times p$  and the expected value  $= \mu$  is of order  $P \times 1$  vector.

Now by considering the partitioning  $X$  into two components  $X_1$  and  $X_2$  of orders  $q$  and  $p - q$

respectively where  $q \leq p$ , that is

$$X = \begin{bmatrix} X^{(1)} \\ X^{(2)} \end{bmatrix}$$

Then we can deduce that

- (a) The marginal distribution of  $X_1$ , and
  - (b) The conditional distribution of  $X_2$  given that  $X_1 = x_1$
- .Both (a) and (b) are also separately multivariate's normal distributions

**Proof:** Since we know that  $\textcolor{brown}{X} \sim N_p(\mu, \Sigma)$ , then Pdf of  $X$  is  $f(X) = \frac{e^{-\frac{1}{2}[X-\mu]\Sigma^{-1}[X-\mu]}}{(2\pi)^{\frac{p}{2}} |\Sigma|^{\frac{1}{2}}}$  and if  $\textcolor{brown}{X}$  is partitioned so that  $X =$

$$\begin{aligned}
& \begin{bmatrix} X_1 \\ X_2 \\ \vdots \\ X_q \\ \dots \\ X_{q+1} \\ \vdots \\ \vdots \\ X_p \end{bmatrix} = \begin{bmatrix} X^{(1)} \\ X^{(2)} \end{bmatrix} = X \quad \text{and} \quad \mu = \begin{bmatrix} \mu^{(1)} \\ \mu^{(2)} \end{bmatrix} \quad \text{and} \quad \Sigma = \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix} f(X^{(1)}, X^{(2)}) = \frac{e^{-\frac{1}{2} [X^{(1)} - \mu^{(1)} X^{(2)} - \mu^{(2)}] \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix}^{-1} [X^{(1)} - \mu^{(1)} X^{(2)} - \mu^{(2)}]}}{(2\pi)^{p/2} |\Sigma_{11}|^{\frac{1}{2}} |\Sigma_{22}|^{\frac{1}{2}} (1 - \rho^2)^{\frac{1}{2}}} \\
& \Sigma^{-1} = \begin{pmatrix} \Sigma_{22} & -\Sigma_{12} \\ -\Sigma_{21} & \Sigma_{11} \end{pmatrix} \begin{array}{c} / \\ \diagdown \end{array} \Sigma_{22} \Sigma_{11} - \Sigma_{12} \Sigma_{21} \quad \Sigma_{22} \Sigma_{11} - \Sigma_{12} \Sigma_{21} = \Sigma_{22} \Sigma_{11} ((1 - \rho^2)^{\frac{1}{2}}) \quad \text{and} \quad f(X^{(1)}) = \int f(X^{(1)}, X^{(2)}) dX^{(2)} = \frac{e^{-\frac{1}{2} [X^{(1)} - \mu^{(1)}] \Sigma_{11}^{-1} [X^{(1)} - \mu^{(1)}]}}{(2\pi)^{q/2} |\Sigma_{11}|^{\frac{1}{2}}} \\
& f(X^{(1)}, X^{(2)}) = \frac{e^{-\frac{1}{2 \Sigma_{11} \Sigma_{11} (1 - \rho^2)} \left[ [X^{(1)} - \mu^{(1)}] \Sigma_{22} [X^{(1)} - \mu^{(1)}] \right] - 2 \left[ [X^{(1)} - \mu^{(1)}] \Sigma_{12} [X^{(2)} - \mu^{(2)}] \right] + \left[ [X^{(2)} - \mu^{(2)}] \Sigma_{11} [X^{(2)} - \mu^{(2)}] \right]}}{(2\pi)^{q/2} (2\pi)^{p-q/2} |\Sigma_{11}|^{\frac{1}{2}} |\Sigma_{22}|^{\frac{1}{2}} (1 - \rho^2)^{\frac{1}{2}}} \quad \text{and} \\
& f(X^{(1)}, X^{(2)}) = \frac{e^{-\frac{1}{2(1-\rho^2)} \left[ [X^{(1)} - \mu^{(1)}] \Sigma_{11}^{-1} [X^{(1)} - \mu^{(1)}] \right] - 2 \left[ [X^{(1)} - \mu^{(1)}] \Sigma_{12} \cancel{\Sigma_{11} \Sigma_{22}} [X^{(2)} - \mu^{(2)}] \right] + \left[ [X^{(2)} - \mu^{(2)}] \Sigma_{11}^{-1} \cancel{\Sigma_{22}} [X^{(2)} - \mu^{(2)}] \right]}}{(2\pi)^{q/2} (2\pi)^{p-q/2} |\Sigma_{11}|^{\frac{1}{2}} |\Sigma_{22}|^{\frac{1}{2}} (1 - \rho^2)^{\frac{1}{2}}} \\
& f(X^{(1)}, X^{(2)}) = \frac{e^{-\frac{1}{2(1-\rho^2)} \left[ [X^{(1)} - \mu^{(1)}] \Sigma_{11}^{-1} [X^{(1)} - \mu^{(1)}] \right] - 2 \left[ [X^{(1)} - \mu^{(1)}] \Sigma_{12} \cancel{\Sigma_{11} \Sigma_{22}} [X^{(2)} - \mu^{(2)}] \right] + \left[ [X^{(2)} - \mu^{(2)}] \Sigma_{11}^{-1} \cancel{\Sigma_{22}} [X^{(2)} - \mu^{(2)}] \right] + \left[ (X_1 - \mu_1) \rho^2 \Sigma_{11}^{-1} ((X_1 - \mu_1)^\top - (X_1 - \mu_1)) \rho^2 \Sigma_{11}^{-1} ((X_1 - \mu_1)^\top) \right]}}{(2\pi)^{q/2} (2\pi)^{p-q/2} |\Sigma_{11}|^{\frac{1}{2}} |\Sigma_{22}|^{\frac{1}{2}} (1 - \rho^2)^{\frac{1}{2}}} \\
& f(X^{(1)}, X^{(2)}) = \frac{e^{-\frac{1}{2(1-\rho^2) \Sigma_{22}} \left[ [X^{(2)} - \mu^{(2)}] \rho \Sigma^{-\frac{1}{2}}_{22} \Sigma^{-\frac{1}{2}}_{11} (X^{(1)} - \mu^{(1)}) \right] \left[ [X^{(2)} - \mu^{(2)}] \rho \Sigma^{-\frac{1}{2}}_{22} \Sigma^{-\frac{1}{2}}_{11} (X^{(1)} - \mu^{(1)}) \right]^\top \Sigma_{11}^{-1} [X^{(1)} - \mu^{(1)}]}}{(2\pi)^{q/2} (2\pi)^{p-q/2} |\Sigma_{11}|^{\frac{1}{2}} |\Sigma_{22}|^{\frac{1}{2}} (1 - \rho^2)^{\frac{1}{2}}} e^{\frac{-1}{2} [X^{(1)} - \mu^{(1)}] \Sigma_{11}^{-1} [X^{(1)} - \mu^{(1)}]^\top} \\
& f(X^{(1)}, X^{(2)}) = \frac{(X_1 - \mu_1) \rho^2 \Sigma_{11}^{-1} ((X_1 - \mu_1)^\top)}{(2\pi)^{q/2} (2\pi)^{p-q/2} |\Sigma_{11}|^{\frac{1}{2}} |\Sigma_{22}|^{\frac{1}{2}} (1 - \rho^2)^{\frac{1}{2}}} \\
& f(X^{(1)}, X^{(2)}) = \frac{e^{-\frac{1}{2(1-\rho^2) \Sigma_{22}} \left[ [X^{(2)} - \mu^{(2)}] \rho \Sigma^{-\frac{1}{2}}_{22} \Sigma^{-\frac{1}{2}}_{11} (X^{(1)} - \mu^{(1)}) \right] \left[ [X^{(2)} - \mu^{(2)}] \rho \Sigma^{-\frac{1}{2}}_{22} \Sigma^{-\frac{1}{2}}_{11} (X^{(1)} - \mu^{(1)}) \right]^\top \Sigma_{11}^{-1} [X^{(1)} - \mu^{(1)}]}}{(2\pi)^{q/2} (2\pi)^{p-q/2} |\Sigma_{11}|^{\frac{1}{2}} |\Sigma_{22}|^{\frac{1}{2}} (1 - \rho^2)^{\frac{1}{2}}} e^{\frac{-1}{2} [X^{(1)} - \mu^{(1)}] \Sigma_{11}^{-1} [X^{(1)} - \mu^{(1)}]^\top} \\
& f(X^{(1)}, X^{(2)}) = \frac{e^{-\frac{1}{2(1-\rho^2) \Sigma_{22}} \left[ [X^{(2)} - \mu^{(2)}] \rho \Sigma^{-\frac{1}{2}}_{22} \Sigma^{-\frac{1}{2}}_{11} (X^{(1)} - \mu^{(1)}) \right] \left[ [X^{(2)} - \mu^{(2)}] \rho \Sigma^{-\frac{1}{2}}_{22} \Sigma^{-\frac{1}{2}}_{11} (X^{(1)} - \mu^{(1)}) \right]^\top \Sigma_{11}^{-1} [X^{(1)} - \mu^{(1)}]}}{(2\pi)^{q/2} (2\pi)^{p-q/2} |\Sigma_{11}|^{\frac{1}{2}} |\Sigma_{22}|^{\frac{1}{2}} (1 - \rho^2)^{\frac{1}{2}}} \\
& f(X^{(2)} \cancel{/} X^{(1)}) = \frac{e^{-\frac{1}{2(1-\rho^2) \Sigma_{22}} \left[ [X^{(2)} - \mu^{(2)}] \rho \Sigma^{-\frac{1}{2}}_{22} \Sigma^{-\frac{1}{2}}_{11} (X^{(1)} - \mu^{(1)}) \right] \left[ [X^{(2)} - \mu^{(2)}] \rho \Sigma^{-\frac{1}{2}}_{22} \Sigma^{-\frac{1}{2}}_{11} (X^{(1)} - \mu^{(1)}) \right]^\top \Sigma_{11}^{-1} [X^{(1)} - \mu^{(1)}]}}{(2\pi)^{q/2} (2\pi)^{p-q/2} |\Sigma_{11}|^{\frac{1}{2}} |\Sigma_{22}|^{\frac{1}{2}} (1 - \rho^2)^{\frac{1}{2}}}
\end{aligned}$$

The marginal distributions of  $X_1$  and  $X_2$  are also normal with mean vector  $\mu_i$  and covariance matrix  $(\Sigma_{ii}, i=1, 2)$ , respectively.

Then

of order qx1, (p-q)x1 i.e. with a similar partition of  $\mu$  and  $\Sigma$

$$\mu = \begin{bmatrix} \mu^{(1)} \\ \mu^{(2)} \end{bmatrix} \quad \text{and} \quad \Sigma = \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix} \quad \text{of order ppx where}$$

**The Joint Probability Distribution of  $f(X)$  is**

$$f(X^{(1)}, X^{(2)}) = \frac{e^{-\frac{1}{2} [X^{(1)} - \mu^{(1)} X^{(2)} - \mu^{(2)}] \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix}^{-1} [X^{(1)} - \mu^{(1)} X^{(2)} - \mu^{(2)}]}}{(2\pi)^{p/2} |\Sigma_{11}|^{\frac{1}{2}} |\Sigma_{22}|^{\frac{1}{2}} (1 - \rho^2)^{\frac{1}{2}}}$$

Where  $\Sigma^{-1} = \begin{pmatrix} \Sigma_{22} & -\Sigma_{12} \\ -\Sigma_{21} & \Sigma_{11} \end{pmatrix}$   $\frac{\Sigma_{22}\Sigma_{11} - \Sigma_{12}\Sigma_{21}}{\Sigma_{22}\Sigma_{11} - \Sigma_{12}\Sigma_{21}} = \Sigma_{22}\Sigma_{11} - \Sigma_{12}\Sigma_{21} = \Sigma_{22}\Sigma_{11}(1-\rho^2)^{\frac{1}{2}}$  and marginal probability of  $X^{(1)}$  is

$$f(X^{(1)}) = \int f(X^{(1)}, X^{(2)}) dX^{(2)} = \frac{e^{-\frac{1}{2}[X^{(1)} - \mu^{(1)}]\Sigma_{11}^{-1}[X^{(1)} - \mu^{(1)}]}}{(2\pi)^{\frac{q}{2}} |\Sigma_{11}|^{\frac{1}{2}}} \quad \text{where}$$

$$f(X^{(1)}, X^{(2)}) = \frac{e^{-\frac{1}{2\Sigma_{11}\Sigma_{11}(1-\rho^2)}[X^{(1)} - \mu^{(1)}][X^{(1)} - \mu^{(1)}] - 2[X^{(1)} - \mu^{(1)}][X^{(2)} - \mu^{(2)}] + [X^{(2)} - \mu^{(2)}][X^{(2)} - \mu^{(2)}]}}{(2\pi)^{\frac{q}{2}} (2\pi)^{\frac{p-q}{2}} |\Sigma_{11}|^{\frac{1}{2}} |\Sigma_{22}|^{\frac{1}{2}} (1-\rho^2)^{\frac{1}{2}}} \quad \text{by simplifying}$$

we get

$$f(X^{(1)}, X^{(2)}) = \frac{e^{-\frac{1}{2(1-\rho^2)}[X^{(1)} - \mu^{(1)}][X^{(1)} - \mu^{(1)}] - 2[X^{(1)} - \mu^{(1)}][X^{(2)} - \mu^{(2)}] + [X^{(2)} - \mu^{(2)}][X^{(2)} - \mu^{(2)}]}}{(2\pi)^{\frac{q}{2}} (2\pi)^{\frac{p-q}{2}} |\Sigma_{11}|^{\frac{1}{2}} |\Sigma_{22}|^{\frac{1}{2}} (1-\rho^2)^{\frac{1}{2}}}$$

for completing square we add and subtract  $(X_1 - \mu_1)\rho^2\Sigma_{11}^{-1}(X_1 - \mu_1)$

$$f(X^{(1)}, X^{(2)}) = \frac{e^{-\frac{1}{2(1-\rho^2)}[X^{(1)} - \mu^{(1)}][X^{(1)} - \mu^{(1)}] - 2[X^{(1)} - \mu^{(1)}][X^{(2)} - \mu^{(2)}] + [X^{(2)} - \mu^{(2)}][X^{(2)} - \mu^{(2)}] + [(X_1 - \mu_1)\rho^2\Sigma_{11}^{-1}(X_1 - \mu_1) - (X_1 - \mu_1)\rho^2\Sigma_{11}^{-1}(X_1 - \mu_1)]}}{(2\pi)^{\frac{q}{2}} (2\pi)^{\frac{p-q}{2}} |\Sigma_{11}|^{\frac{1}{2}} |\Sigma_{22}|^{\frac{1}{2}} (1-\rho^2)^{\frac{1}{2}}}$$

By simplifying, we get

$$f(X^{(1)}, X^{(2)}) = \frac{e^{-\frac{1}{2(1-\rho^2)\Sigma_{22}}[X^{(2)} - \mu^{(2)} - \rho\Sigma_{22}^{-\frac{1}{2}}\Sigma_{11}^{-\frac{1}{2}}(X^{(1)} - \mu^{(1)})][X^{(2)} - \mu^{(2)} - \rho\Sigma_{22}^{-\frac{1}{2}}\Sigma_{11}^{-\frac{1}{2}}(X^{(1)} - \mu^{(1)})]}}{(2\pi)^{\frac{q}{2}} (2\pi)^{\frac{p-q}{2}} |\Sigma_{11}|^{\frac{1}{2}} |\Sigma_{22}|^{\frac{1}{2}} (1-\rho^2)^{\frac{1}{2}}} e^{\frac{-1}{2}[X^{(1)} - \mu^{(1)}]\Sigma_{11}^{-1}[X^{(1)} - \mu^{(1)}]}$$

$$\text{As we know that } f(x_2/x_1) = \frac{f(X^{(1)}, X^{(2)})}{f(X^{(1)})} =$$

$$\frac{e^{-\frac{1}{2(1-\rho^2)\Sigma_{22}}[X^{(2)} - \mu^{(2)} - \rho\Sigma_{22}^{-\frac{1}{2}}\Sigma_{11}^{-\frac{1}{2}}(X^{(1)} - \mu^{(1)})][X^{(2)} - \mu^{(2)} - \rho\Sigma_{22}^{-\frac{1}{2}}\Sigma_{11}^{-\frac{1}{2}}(X^{(1)} - \mu^{(1)})]}}{(2\pi)^{\frac{q}{2}} (2\pi)^{\frac{p-q}{2}} |\Sigma_{11}|^{\frac{1}{2}} |\Sigma_{22}|^{\frac{1}{2}} (1-\rho^2)^{\frac{1}{2}}} e^{\frac{-1}{2}[X^{(1)} - \mu^{(1)}]\Sigma_{11}^{-1}[X^{(1)} - \mu^{(1)}]}$$

$$f(X^{(2)} / X^{(1)}) = \frac{e^{-\frac{1}{2(1-\rho^2)\Sigma_{22}}[X^{(2)} - \mu^{(2)} - \rho\Sigma_{22}^{-\frac{1}{2}}\Sigma_{11}^{-\frac{1}{2}}(X^{(1)} - \mu^{(1)})][X^{(2)} - \mu^{(2)} - \rho\Sigma_{22}^{-\frac{1}{2}}\Sigma_{11}^{-\frac{1}{2}}(X^{(1)} - \mu^{(1)})]}}{(2\pi)^{\frac{p-q}{2}} |\Sigma_{22}|^{\frac{1}{2}} (1-\rho^2)^{\frac{1}{2}}}$$

$$\frac{X_2}{X_1} \sim N_{p-q} \left( \mu_2 + \rho \Sigma_{22}^{-\frac{1}{2}} \Sigma_{11}^{-\frac{1}{2}} (X_1 - \mu_1), (1-\rho^2) \Sigma_{22} \right)$$

$$\frac{X_2}{X_1} \sim N_{p-q} \left( \mu_2 + \frac{\Sigma_{12}\Sigma_{22}^{-\frac{1}{2}}\Sigma_{11}^{-\frac{1}{2}}}{\Sigma_{11}^{\frac{1}{2}}\Sigma_{22}^{\frac{1}{2}}} (X_1 - \mu_1), (1 - \frac{\Sigma_{12}\Sigma_{21}}{\Sigma_{11}\Sigma_{22}}) \Sigma_{22} \right)$$

$$\frac{X_2}{X_1} \sim N_{p-q} [(\mu_2 + \Sigma_{12}\Sigma_{11}^{-1}(X_1 - \mu_1), (\Sigma_{22} - \Sigma_{12}\Sigma_{11}^{-1}\Sigma_{21})]$$

Which fulfill the requirement of multivariate normal distribution.

These result can be verified for if  $X_1$  and  $X_2$  are independent i.e.  $\rho=0$  ( $\Sigma_{12}=0$ ) and  $X_2/X_1$  are independent.

$$X_2/X_1 = X_1 \sim N_1(\mu_2, \Sigma_{22})$$

This result also verify if  $X$  is bidimensional random variable then the resulting distribution is bivariate conditional normal distribution i.e. if  $q=1$  and  $p=2$  then  $\Sigma_{11}=\sigma_1^2$ ,  $\Sigma_{12}=\sigma_{12}$  and  $\Sigma_{22}=\sigma_2^2$ .

$X_2/X_1 = X_1 \sim N_1[(\mu_2 + \sigma_{12}(X_1 - \mu_1)/\sigma_1^2, \sigma_2^2 - (\sigma_{12}\sigma_{21})/\sigma_1^2]$  and again if  $X_2$  is independent of  $X_1$ , then  $\sigma_{12}=\sigma_{21}=0$  and the result fulfill univariate normal distribution

$$X_2/X_1 = X_1 \sim N_1[\mu_2, \sigma_2^2]$$

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